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| Ordered SCD: A New Stochastic Optimization Framework | | | | |

for Empirical Risk Minimization (AISTATS 2020)

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Optimization Problem

Consider the Optimization problem for minimizing the average of loss function with regularizer, let $L(\theta)$,

$$\min_{\theta} \frac{1}{n} \sum_{i=1}^{n} L_i(\theta) + R(\theta)$$

where θ is the parameter vector of the parameterized model and $L_i's$ are loss of the *i*-th data sample, and $R(\theta) \ge 0$ is regularizer.

• We can use Gradient based methods which iteratively update the parameters as follows

$$\theta_{k+1} = \theta_k - \alpha \nabla L(\theta_k)$$

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| Gradient Desc | rent | | | |

 As we aleady know, the Gradient Descent method updates the parameter by using all the gradients of dataset

1 2 3 4 Dataset of size n n-1 n
$$\theta_{k+1} = \theta_k - \alpha \nabla \frac{1}{n} \sum_{i=1}^n L_i(\theta_k)$$

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| Gradient D | escent | | | |

As we aleady know, the Gradient Descent method updates the parameter by using all the gradients of dataset. i.e. full gradient



Computational expensive, inefficient

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Stochastic Gradient Descent

 To overcome the inefficiency, we devide the dataset to mini-batch and we replace the full gradient with mini-batch gradient(gradient estimator)



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Stochastic Gradient Descent

Is this a reasonable method?



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Stochastic Gradient Descent

Sure



SGD is a unbiased method

• Quite efficient and sometimes it is better than GD for use in DNN, which is non-convex.

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| Intuition | | | | |

• When determining a decision boundary, there seem to be more impactful, helpful samples



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Intuition



- Is uniform sampling always better?
- If not, how can we get better?
- Loss values indicate how violate from the answer.

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| Intuition | | | | |



• Let's use only top-q loss valued samples within a batch, instead of using all of a batch

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Algorithm

Definition 1

Given a set of n real numbers (a_1, a_2, \cdots, a_n) , an index subset $S \subseteq \{1, 2, \cdots, n\}$, and a positive integer number $q \leq |S|$, we define q-argmax_{$j \in S$} a_j such that $Q \in q$ -argmax_{$j \in S$} a_j is a set of q indices of the q-largest values of $(a_j)_{j \in S}$; i.e., q-argmax_{$j \in S$} a_j = argmax_{$Q \subseteq S, |Q| = q$} $\sum_{i \in Q} a_i$

• For example, let $(a_1 = -5, a_2 = 10, a_3 = -4, a_4 = 6, a_5 = -1, a_6 = 5)$, if we want to know the 3 highest value indices of index subset $S = \{1, 2, 3, 4, 5\}$, find the 3 - argmaxset_{$j \in S$} $a_j = \{2, 4, 5\}$ In our case, It returns the top-q largest loss valued data indices. Then we can use only these data samples.

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| Algorithm | | | | |

Algorithm 1 Ordered Stochastic Gradient Descent(OSGD)

Input : problem data L(x), step sizes sequence $(\alpha_k)_{k \in \mathbb{N} \cup \{0\}}$ and initialization θ_0

for k = 0, 1, ... do Sample a mini-batch uniformly: $S \subseteq \{1, 2, ..., n\}$ with |S| = sFind a set Q_k of top-q samples in S in term of loss values: $Q_k \in q$ -argmax $_{t \in S} L_t(\theta_k)$ Compute a gradient $\tilde{g}_k = \nabla_{\theta} L_{Q_k}(x_k)$ where $L_{\theta_k}(x_k) = \frac{1}{q} \sum_{t \in Q_k} L_t(\theta_k)$ Update parameter $\theta_{k+1} = \theta_k - \alpha_k \tilde{g}_k$ end for

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• Notice that $\nabla_{\theta} L_{Q_k}$ is a biased gradient estimator

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Toy examples



Figure 1: Decision boundaries of mini-batch SGD predictors (**top** row) and ordered SGD predictors (**bottom** row) with 2D synthetic datasets for binary classification. In these examples, ordered SGD predictors correctly classify more data points than mini-batch SGD predictors, because a ordered SGD predictor can focus more on a smaller yet informative subset of data points, instead of focusing on the average loss dominated by a larger subset of data points.

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Ordering notation

 Let's bring a notation similar to order statistics, but descending order. the notation of ordered indices : given a model parameter θ, let

 $L_{(1)}(\theta) \ge L_{(2)}(\theta) \ge \cdots \ge L_{(n)}(\theta)$

be the decreasing values of the individual losses $L_1(\theta), \dots, L_n(\theta)$, where $(j) \in \{1, \dots, n\}$. That is, $\{(1), \dots, (n)\}$ as a permutation of $\{1, \dots, n\}$ defines the order of sample indices by loss values.

• For example, let $L_1(\theta) = 1, L_2(\theta) = 10, L_3(\theta) = 2, L_4(\theta) = 6, L_5(\theta) = 3, L_6(\theta) = 5$, then we get the ordering notation by given θ , $L_{(6)}(\theta) = 1, L_{(1)}(\theta) = 10, L_{(5)}(\theta) = 2, L_{(2)}(\theta) = 6, L_{(4)}(\theta) = 3, L_{(3)}(\theta) = 5$

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What we actually optimized is

Theorem 1

Consider the following objective function:

$$L_q(\theta) := rac{1}{q} \sum_{t=1}^n \gamma_t L_{(t)}(\theta)$$

b where the parameter γ_t depends on the hyper parameter tuple (n,s,q), and is defined by

$$\gamma_t := \frac{\sum_{l=0}^{q-1} {\binom{t-1}{l} \binom{n-t}{s-l-1}}}{\binom{n}{s}}$$

Then, Ordered Stochastic Gradient Descent is a stochastic first-order method for minimizing $L_q(x)$ in sense that \tilde{g}_k is used in OSGD is an unbiased estimator of a gradient of $L_q(x)$.

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Example

•
$$L_q(\theta) := \frac{1}{q} \sum_{t=1}^n \gamma_t L_{(t)}(\theta)$$
 where $\gamma_t := \frac{\sum_{l=0}^{q-1} {t-1 \choose l} {n-t \choose s-l-1}}{{n \choose s}}$

For the case of
$$(n=12, s=4, q=2)$$
,

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Take a close look

$$L_q(\theta) := \frac{1}{q} \sum_{t=1}^n \gamma_t L_{(t)}(\theta) = \sum_{t=1}^n \frac{\gamma_t}{q} L_{(t)}(\theta)$$

• For the case of (n=12, s=4, q=2),

$$\frac{1}{q}\gamma_1 = 0.1667 \quad \frac{1}{q}\gamma_2 = 0.1667 \quad \frac{1}{q}\gamma_3 = 0.1576 \quad \frac{1}{q}\gamma_4 = 0.1414 \quad \frac{1}{q}\gamma_5 = 0.1202 \quad \frac{1}{q}\gamma_6 = 0.0808$$
$$\frac{1}{q}\gamma_7 = 0.0707 \quad \frac{1}{q}\gamma_8 = 0.0464 \quad \frac{1}{q}\gamma_9 = 0.0253 \quad \frac{1}{q}\gamma_{10} = 0.0091 \quad \frac{1}{q}\gamma_{11} = 0 \quad \frac{1}{q}\gamma_{12} = 0$$
$$\sum_{t=1}^n \frac{\gamma_t}{q} = 1$$

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Asymptotic behavior of the γ_t

Proposition 1

Denote
$$z = \frac{j}{n}$$
 and $\gamma(z) := \sum_{l=0}^{q-1} z^l (1-z)^{s-l-1} \frac{s!}{l!(s-l-1)!}$. Then it holds that
$$\lim_{j,n\to\infty} \gamma_j = \frac{1}{n} \gamma(z)$$

Moreover, it holds that
$$1 - \frac{1}{s}\gamma(z)$$
 is the cumulative distribution function of $Beta(z;q,s-q)$.

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Asymptotic behavior of the γ_t



Figure 2: $\hat{\gamma}(z)$ and $\gamma(z)$ for different (n, s, q) where $\hat{\gamma}$ is a rescaled version of γ_j : $\hat{\gamma}(j/n) = n\gamma_j$.

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Convergence Analysis

Theorem 2

Let $\{\theta_t\}_{t=0}^T$ be a sequence generated by ordered SGD (Algorithm1). Suppose that $L(\cdot)$ is G_1 -Lipschitz continuous for $i = 1, \cdots, n$, and $R(\cdot)$ is G_2 -Lipschitz continuous. Suppose that there exists a finite $\theta_{\star} \in \operatorname{argmin}_{\theta} L_q(\theta)$ and $L_q(\theta_{\star})$ is finite. Then, the following two statements hold:

1 (Convex setting). If $L_i(\cdot)$ and $R(\cdot)$ are both convex, for any step-size η_t , it holds that

$$\min_{0 \le t \le n} \mathbb{E}[L_q(\theta_t) - L_q(\theta_\star)] \le \frac{2(G_1^2 + G_2^2) \sum_{t=0}^T \eta_t^2 + ||\theta_\star - \theta_0||^2}{2 \sum_{t=0}^T \eta_t}$$

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Convergence Analysis

Theorem 2

Let $\{\theta_t\}_{t=0}^T$ be a sequence generated by ordered SGD [Algorithm1]. Suppose that $L(\cdot)$ is G_1 -Lipschitz continuous for $i = 1, \cdots, n$, and $R(\cdot)$ is G_2 -Lipschitz continuous. Suppose that there exists a finite $\theta_{\star} \in \operatorname{argmin}_{\theta} L_q(\theta)$ and $L_q(\theta_{\star})$ is finite. Then, the following two statements hold:

1 (Convex setting). If $L_i(\cdot)$ and $R(\cdot)$ are both convex, for any step-size η_t , it holds that

if we choose
$$\eta_t \sim \mathcal{O}(\frac{1}{\sqrt{t}})$$
,
the optimality gap $\min_{0 \le t \le n} (L_q(\theta_t) - L_q(\theta_\star))$ decay at the rate of $\tilde{\mathcal{O}}(\frac{1}{\sqrt{t}})$

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Hyper-parameter setting

Basically, OSGD uses the adaptive q-value setting which the default setting: q=s at the beginning of training, $q = \lfloor \frac{s}{2} \rfloor$ once train $acc \ge 80\%$, $q = \lfloor \frac{s}{4} \rfloor$ once train $acc \ge 90\%$, $q = \lfloor \frac{s}{8} \rfloor$ once train $acc \ge 95\%$, and $q \lfloor \frac{s}{16} \rfloor$ once train $acc \ge 99.5\%$, where train acc represents training accuracy.

This rule was derived based on the intuition that in the early stage of training, all samples are informative to build a rough model, while the samples around the boundary (with larger losses) are more helpful to build the final classifier in later stage.

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Experimental results



Figure 3: Test error and training loss (in log scales) versus the number of epoch. These are without data augmentation in subfigures (a)-(d), and with data augmentation in subfigures (e)-(h). The lines indicate the mean values over 10 random trials, and the shaded regions represent intervals of the sample standard deviations.

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Take a close look



- Model = PreActResNet18
- (n,s,q) = (50000, 64, 64 or 30 or adaptive)
- \blacksquare the initial LR = 0.01 and decay with some policy
- 10 trial each and averaged

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Wall-clock time

| Data Aug | Datasets | Model | mini-batch SGD | ordered SGD | difference |
|----------|---------------|---------------------------------|----------------|--------------|------------|
| No | Semeion | Logistic model | 0.15(0.01) | 0.15(0.01) | 0.00 |
| No | MNIST | Logistic model | 7.16(0.27) | 7.32(0.24) | -0.16 |
| No | Semeion | $_{\rm SVM}$ | 0.17(0.01) | 0.17(0.01) | 0.00 |
| No | MNIST | $_{\rm SVM}$ | 8.60(0.31) | 8.72(0.29) | -0.12 |
| No | Semeion | LeNet | 0.18(0.01) | 0.18(0.01) | 0.00 |
| No | MNIST | LeNet | 9.00(0.34) | 9.12(0.27) | -0.12 |
| No | KMNIST | LeNet | 9.23(0.33) | 9.04~(0.55) | 0.19 |
| No | Fashion-MNIST | LeNet | 8.56(0.48) | 9.45~(0.31) | -0.90 |
| No | CIFAR-10 | $\operatorname{PreActResNet18}$ | 45.55(0.47) | 43.72(0.93) | 1.82 |
| No | CIFAR-100 | $\operatorname{PreActResNet18}$ | 46.83(0.90) | 43.95(1.03) | 2.89 |
| No | SVHN | $\operatorname{PreActResNet18}$ | 71.95(1.40) | 66.94(1.67) | 5.01 |
| Yes | Semeion | LeNet | 0.28(0.02) | 0.28(0.02) | 0.00 |
| Yes | MNIST | LeNet | 14.44 (0.54) | 14.77(0.41) | -0.32 |
| Yes | KMNIST | LeNet | 12.17(0.33) | 11.42(0.29) | 0.75 |
| Yes | Fashion-MNIST | LeNet | 12.23(0.40) | 12.38(0.37) | -0.14 |
| Yes | CIFAR-10 | $\operatorname{PreActResNet18}$ | 48.18(0.58) | 46.40(0.97) | 1.78 |
| Yes | CIFAR-100 | $\operatorname{PreActResNet18}$ | 47.37 (0.84) | 44.74 (0.91) | 2.63 |
| Yes | SVHN | PreActResNet18 | 72.29 (1.23) | 67.95(1.54) | 4.34 |

Table 4: Average wall-clock time (seconds) per epoch.

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Varying q-size



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 This purposely biased gradient estimator perform well not only empirical risk minimization but also perspective of generalization and computational efficiency

This variant of SGD also guaranteed to converge

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Thanks for your attention

proof of Theorem 1

Need to find the function that the \tilde{g}_k become an unbiased estimator for subgradient Taking expectation to \tilde{g}_k , it holds that

$$\mathbb{E}[\tilde{g}_k] = \frac{1}{q} \mathbb{E}[\sum_{i \in Q_k} g_i] = \frac{1}{q} \sum_{i=1}^n P(i \in Q_k) g_i = \frac{1}{q} \sum_{j=1}^n P((j) \in Q_k) g_{(j)}$$

Define index set $A_j = \{(1), (2), ..., (j-1)\}$, denote that given order is measured from whole n sample losses at current parameter x_k , then

1

$$\begin{split} P((j) \in Q) &= P((j) \in \operatorname{q-argmax}_{t \in S} L_t(x_k)) \\ &= P((j) \in S \text{ and } S \text{ contains at most q-1 items in } A_j) \\ &= P((j) \in S) P(S \text{ contains at most q-1 items in } A_j | (j) \in S) \\ &= P((j) \in S) \sum_{l=0}^{q-1} P(S \text{ contains } l \text{ items in } A_j | (j) \in S) \end{split}$$

Then there are $\binom{n}{s}$ different sets S s.t |S| = s and $\binom{n-1}{s-1}$ different sets S contains index (j). So $P((j) \in S) = \frac{\binom{n-1}{s-1}}{\binom{n}{s}}$.

And given the condition $(j) \in S$, let S contains l items in A_j which implies s - l - 1 items in $\{(j + 1), (j + 2), ..., (n)\}$. Then it holds that

$$P(S \text{ contains } l \text{ items in } A_j | (j) \in S) = \frac{\binom{j-1}{l}\binom{n-j}{s-l-1}}{\binom{n-1}{s-1}}$$

Therefore

$$P((j) \in Q_k) = \frac{\binom{n-1}{s-1}}{\binom{n}{s}} \sum_{l=0}^{q-1} \frac{\binom{j-1}{l}\binom{n-j}{s-l-1}}{\binom{n-1}{s-1}} = \frac{\sum_{l=0}^{q-1} \binom{j-1}{l}\binom{n-j}{s-l-1}}{\binom{n}{s}} =: \gamma_j$$

So the expectation of ordered gradient is

$$\mathbb{E}[\tilde{g}_k] = \frac{1}{q} \sum_{j=1}^n P((j) \in Q_k) g_{(j)} = \frac{1}{q} \sum_{j=1}^n \gamma_j g_{(j)}$$

which desired. Then the proof is done.